## RESEARCH

**Open Access** 



# Association between diabetes mellitus and multi-drug-resistant tuberculosis: evidence from a systematic review and meta-analysis

Balewgizie Sileshi Tegegne<sup>1,2</sup>, Melkamu Merid Mengesha<sup>1\*</sup>, Andreas A. Teferra<sup>3</sup>, Mamaru Ayenew Awoke<sup>4</sup> and Tesfa Dejenie Habtewold<sup>2</sup>

## Abstract

**Background:** Diabetes mellitus (DM) poses a significant risk for the development of active tuberculosis (TB) and complicates its treatment. However, there is inconclusive evidence on whether the TB-DM co-morbidity is associated with a higher risk of developing multi-drug-resistant tuberculosis (MDR-TB). The aim of this meta-analysis was to summarize available evidence on the association of DM and MDR-TB and to estimate a pooled effect measure.

**Methods:** PubMed, Excerpta Medica Database (EMBASE), Web of Science, World Health Organization (WHO), and Global Health Library database were searched for all studies published in English until July 2018 and that reported the association of DM and MDR-TB among TB patients. To assess study quality, we used the Newcastle-Ottawa Scale for cohort and case-control studies and the Agency for Healthcare Research and Quality tool for cross-sectional studies. We checked the between-study heterogeneity using the Cochrane Q chi-squared statistic and *I*<sup>2</sup> and examined a potential publication bias by visual inspection of the funnel plot and Egger's regression test statistic. The random-effect model was fitted to estimate the summary effects, odds ratios (ORs), and 95% confidence interval (CIs) across studies.

**Results:** This meta-analysis of 24 observational studies from 15 different countries revealed that DM has a significant association with MDR-TB (OR = 1.97, 95% CI = 1.58–2.45,  $l^2$  = 38.2%, *P* value for heterogeneity = 0.031). The significant positive association remained irrespective of country income level, type of DM, how TB or DM was diagnosed, and design of primary studies. A stronger association was noted in a pooled estimate of studies which adjusted for at least one confounding factor, OR = 2.43, 95% CI 1.90 to 3.12. There was no significant publication bias detected.

**Conclusions:** The results suggest that DM can significantly increase the odds of developing MDR-TB. Consequently, a more robust TB treatment and follow-up might be necessary for patients with DM. Efforts to control DM can have a substantial beneficial effect on TB outcomes, particularly in the case of MDR-TB.

Systematic review registration: PROSPERO CRD42016045692.

Keywords: Diabetes mellitus, Tuberculosis, Multi-drug-resistant tuberculosis, Systematic review, Meta-analysis

## Background

The global tuberculosis (TB) burden continues to be a major public health challenge despite efforts to reduce its impact. Globally in 2016, there were an estimated 10.4 million incident cases of TB, equivalent to 140 cases per 100, 000 population [1]. Of these incident cases of TB in 2016, an estimated 1.9 million were attributed to

<sup>1</sup>Department of Epidemiology and Biostatistics, College of Health and Medical Sciences, Haramaya University, Harar, Ethiopia Full list of author information is available at the end of the article undernourishment, 1.0 million to HIV, 0.8 million to smoking, and 0.8 million to diabetes [1]. In the era of the sustainable development goals, post-2015, the "*End TB Strategy*" targets to reduce TB incidence by 80% by year 2030 [2]. However, the global epidemiological and demographic transitions pose significant challenge to TB control programs by changing the relative importance of different risk factors for TB [3, 4].

The global diabetes mellitus (DM) epidemic poses a significant bottleneck to the TB control program [3, 4]. The International Diabetes Federation (IDF) estimated

© The Author(s). 2018 **Open Access** This article is distributed under the terms of the Creative Commons Attribution 4.0 International License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made. The Creative Commons Public Domain Dedication waiver (http://creativecommons.org/publicdomain/zero/1.0/) applies to the data made available in this article, unless otherwise stated.

<sup>\*</sup> Correspondence: melkamumrd@gmail.com

**BMC** 

that, globally in 2013, 382 million adults have diabetes of whom 80% live in low-and middle-income countries. Further increase in the global burden of diabetes is predicted, reaching 592 million by 2035 [5]. People with diabetes, compared to non-diabetic controls, were twoto three-fold more likely to develop TB [6, 7]. In 2013, an estimated 15% of adult cases of TB worldwide were attributed to diabetes, which corresponds to 1 million cases of diabetes-associated TB per year [3]. Impaired immunity in diabetic patients is thought to contribute to the evolution of latent TB infection to active cases. Moreover, people with TB who have DM have a poorer response to treatment than do those without DM, and are therefore at a higher risk of TB treatment failure, death, and relapse after cure [3, 8]. Treatment failure in turn adds another significant challenge to the global TB control program, a drug-resistant TB [1, 9].

Multi-drug-resistant tuberculosis (MDR-TB), resistance to at least isoniazid and rifampicin, results from either primary infection with resistant bacteria or may develop in the course of a patient's treatment [9]. In 2016, there were an estimated 600,000 incident cases of MDR-TB. In the same year, an estimated 4.1% of new cases and 19% of previously treated TB cases had MDR-TB [1]. The emergence of multi-drug resistance across the world poses a global threat as the treatment is difficult, expensive, and a major healthcare cost burden to developing countries [10]. Most cases of MDR-TB arise from a mixture of physician error, inadequate and incomplete treatment, and patient non-compliance during treatment of susceptible TB [11, 12]. Research reports also indicate that patients with MDR-TB and a co-morbidity of DM have a poor treatment response compared with non-diabetic MDR-TB controls [13].

The additional risk of DM for the development of MDR-TB, however, remains controversial [14–16]. Many previous studies have found a 2.1 to 8.8 times increased risk of MDR-TB among TB patients co-morbid with diabetes [17-21]. In addition, observational studies from Israel, Georgia, and Mexico showed that TB patients with DM had a higher risk of developing MDR-TB [22-24]. In contrast, several others reported that there is no increased risk of MDR-TB among TB patients who have DM [25-28]. Similarly, none but one of the previously conducted systematic reviews and meta-analysis [29] reported DM as an independent risk factor for MDR-TB. However, the pooled estimate in that study was based on limited number of studies which mostly implemented a cross-sectional or case-control study design. By conducting a comprehensive search until July 2018, we identified more studies and included six new cohort studies [30-35]. Therefore, with the present systematic review and meta-analysis, we aimed to assess the pooled effect estimate of DM on the development of MDR-TB with the careful inclusion of data from appropriately conducted observational studies.

## Methods

## Registration

Our systematic review has been registered with the International Prospective Register of Systematic Reviews (PROSPERO) (http://www.crd.york.ac.uk/prospero/display\_record.asp?ID=CRD42016045692). The protocol has been published elsewhere [36]. This review is written in accordance with the recommendations from the Preferred Reporting Items for Systematic Review and Meta-Analysis (PRISMA) statement guideline [37, 38], and a completed PRISMA checklist has been included (Additional file 1: Table S1).

### **Eligibility criteria**

We included all observational studies (cross-sectional cohort, case-control cohort, and prospective and retrospective cohorts) which reported the association of DM and MDR-TB among TB patients. All eligible studies published in English and prior to July 30, 2018, were included for the review.

## Data source and search strategy

PubMed, Excerpta Medica Database (EMBASE), Web of Science, and WHO Global Health Library databases were searched for all publications. We also searched cross-references of identified articles. In consultation with an experienced medical information specialist, a comprehensive search strategy has been developed (Table 1). Search results were compiled using citation management software (RefWorks 2.0; ProQuest LLC, Bethesda, MD, USA, http://www.refworks.com).

## Study selection

Articles were screened and selected for full-text review if they met the following selection criteria: (1) provided or permitted the computation of an effect estimate of DM on the development of MDR-TB; (2) included TB patients (all type) and defined MDR-TB based on standard protocol; resistance at least to isoniazid and rifampicin [9]; and (3) defined DM based on any of the following: baseline diagnosis by self-report, medical records, laboratory test, or treatment with oral hypoglycemic medications or insulin. We excluded studies for any of the following reasons: citations without abstracts, anonymous reports, duplicate studies, case reports, or studies which did not compare MDR-TB among people with DM to people without DM, and systematic reviews and meta-analysis. Additionally, studies that either did not provide effect estimates in odds ratios, rate ratios, hazard ratios, and relative risks or did not allow for the computation of these values were excluded. Two authors

("Diabetes Mellitus" [mesh] OR diabetes

\*[tiah] OR diabetic \*[tiab] OR T2DM [tiab] OR T1DM [tiab] OR "T2

DM"[tiab] OR "T1 DM"[tiab])

("Tuberculosis

[Mesh] OR

Multidrug-Resistant"

(("Tuberculosis"[Mesh]

OR tubercul\*[tiab] OR

\*[tiab]) AND ("Drug

Resistance, Multiple

[Mesh] OR multidrug

resist\*[tiab] OR multi-

drug resist\*[tiab] OR drug resist\*[tiab] OR

multiresist\*[tiab] OR multi resist\*[tiab])) OR rifampcin resist\*[tiab] OR MDR-TB [tiab])

MDR [tiab] OR

#1 AND #2

tb[tiab] OR antitubercul

number	of identified abstracts	per literatu	ire database			
No. of hits	EMBASE	No. of hits	Web of Science	No. of hits	WHO Global Health Library	No. of hits
612,633	('Diabetes Mellitus'/ exp. OR (diabetes* OR diabetic* OR T2DM OR T1DM OR 'T2 DM' OR 'T1 DM'): ab,ti)	1,006,444	TS = (diabetes* OR diabetic* OR T2DM OR T1DM OR "T2 DM" OR "T1 DM")	649,798	((Diabetes Mellitus) OR diabetes* OR diabetic* OR (T2DM) OR (T1DM) OR (T2 DM) OR (T1 DM))	670,864

17.917

254\*

((Tuberculosis

Multidrug-Resistant)

OR ((Tuberculosis)

tb OR antitubercul\*

resist\* OR multi-drug resist\* OR drug resist\*

multiresist\* OR multi

rifampcin resist\*

OR tubercul\* OR

AND (multidrug

OR MDR OR

resist\*) OR

OR MDRTB))

#1 AND #2

TS = (tubercul\* OR

tb OR antitubercul\*)

AND TS = (multidrug

resist\* OR drug

resist\* OR MDR

multi resist\* OR

MDRTB)

#1 AND #2

OR multiresist\* OR

rifampcin resist\* OR

20.105

525\*

Table 1 Search strings used and number

('Tuberculosis

Multidrug-Resistant'/

exp. OR (tubercul\*

resistance'/exp. OR

OR multiresist\* OR

'multi resist\*'):ab,ti))

OR MDRTB):ab.ti)

#1 AND #2

OR ('rifampcin resist\*'

exp. OR (('tuberculosis'/

OR tb OR antitubercul\*):

ab,ti) AND ('multidrug

('multidrug resist\*' OR

'drug resist\*' OR MDR

16.247

\*Date of hits: July 30, 2018

Combined

search

Component PubMed

Diabetes

mellitus

Multi-drug-

tuberculosis

resistant

(BS and MM) screened and checked full-text studies for inclusion independently. Any disagreement was resolved by discussion. If consensus could not be reached, a third author determined the eligibility and approved the final list of retained studies.

235\*

## Quality assessment and data extraction

Meta-analysis of observational studies present particular challenges because of potential biases in the original studies and differences in study designs that make the calculation of a single summary estimate of effect of exposure potentially misleading [39]. Thus, assessing quality of studies using a standardized tool helps to classify risk of bias which can help to explain variation in the results of included studies. Two authors (BS and MM) checked the quality of studies independently using Newcastle-Ottawa Scale (NOS) [40] for cohort and case-control studies and the Agency for Healthcare Research and Quality (AHRQ) [41] tool for cross-sectional studies as shown in Additional file 2: Table S2. Case-control and cohort studies gualified for inclusion if they scored 7 points or more from a total of 9 points in three domains of the equally weighted nine NOS components: selection (4 points), comparability (2 points), and exposure assessment (3 points). Cross-sectional studies were included in the analysis if they fulfilled all the four components (comparability, exposure, outcome measurement, and statistical analysis) of the AHRO criteria. Structured data extraction form was constructed and pre-tested. For every study that met our eligibility criteria, two authors (BS and MM) independently extracted the title, name of authors, year of publication, country, study design, study population, sample size, data collection procedure, diagnosis of DM, and MDR-TB. Crude or adjusted effect sizes (ORs) with confidence intervals in the original studies were also extracted.

## Statistical analysis

We estimated pooled OR with their 95% CI to evaluate the association between DM and MDR-TB among TB patients. Potential sources of heterogeneity between the studies were examined by using the Cochrane Q chi-squared statistic and  $I^2$  [42, 43], where  $I^2 > 75\%$ suggested considerable heterogeneity. Presuming the variation of the true effect of DM on MDR-TB between studies, the random-effect model [44] was fitted to estimate the summary effect (ORs) and 95% CIs across studies. Subgroup analyses were performed by study type, variable adjustment, DM type, and TB type and by the income level of the country where the primary study was conducted. Publication bias was assessed with the funnel plot for asymmetry, Egger's test, and Begger's regression models [45]. To see the trend of evidence

52110

768\*

accumulation, we ran a cumulative meta-analysis. We conducted an influence analysis to observe the effect of omitting a single study on the overall pooled effect estimate [46]. All analyses were performed using STATA SE 14.2 (Stata Corporation, College Station, TX) [47]. All reported *P* values were two-sided at the level of 0.05.

## Results

## Search results

We identified a total of 1782 studies based on the literature search in four databases including PubMed (235), EMBASE (525), Web of Science (254), and WHO Global Health library (768). Additionally, we found nine studies through a manual search (Fig. 1). After removal of duplicate studies, 1112 articles were screened based on titles and abstracts leaving 73 studies selected for a full-text review. Furthermore, 16 records were excluded from the full-text review (7 conference abstracts and correspondence/short communications and 9 full texts not accessible or available). Similarly, from the full-text review, we excluded 36 studies due to different reasons (6 were not written in English, 10 did not include comparisons, 17 did not define outcome clearly, and 3 had no enough outcome to estimate risks). List of the excluded articles after full-text review is available in Additional file 3. Finally, 21 articles were included for data synthesis. Additionally, three articles [31, 48, 49] reported separate effect sizes for newly diagnosed and previously treated TB patients, and one study [50] reported separate effect estimate by country (USA and Mexico), which resulted in a total of 25 studies or data points for analyses.

## **Study characteristics**

Table 2 presents the characteristics of all the 25 studies [14, 18–21, 23, 25, 28, 30–35, 48–54] with a total sample of 13,403 participants with TB. The studies were published between 2001 and 2018 and covered different geographical regions: 13 studies were in Asia (three in China, three in Taiwan, two in Georgia, one in Bangladesh, one in Indonesia, one in Iran, one in Thailand, and one in South Korea), three studies were in Europe (one in Spain, one in Portugal, and one in Turkey), eight studies were in the Americas (four in Mexico, two in the USA, and two in Peru), and one study was in Africa (Egypt). From the total studies included in the analysis, nine were case-control [18–21, 23, 25, 32, 52, 54], eight



Table 2 Charac	teristics of s	studies ir	ncluded in th	ne meta-analysis a:	ssessing the association	ion between diá	abetes mellitus and r	multi-drug	resistant tube	erculosis	
Author	County	Design	TB-DM) (TB-DM)	Data collection procedure	DM diagnosis	Diagnosis of TB	Diagnosis of MDR-TB	TB type	DM type	OR (95% CI) <sup>††</sup>	Adjustment to confounders
Baghaei P et al. [25]	Iran	2	282 (24)	Record review	Not specified	Not specified	Not specified	All	Type not mentioned	0.68 (0.19–2.36)	No
Bashar M et al. [18]	USA	2	155 (50)	Record review	Not specified	Not specified	Not specified	AII	T1 and T2	5.30 (1.90–14.70)	Homelessness and HIV status
Carreira S et al. [34]	Portugal	RC	246 (123)	Record review	FBS/previous history of DM and receiving insulin or oral hypoglycemic agents	Not specified	Not specified	AII	Type not mentioned	1.49 (0.24–9.06)	oN
Chang JT [14]	Taiwan	D	192 (60)	Record review/ clinical/laboratory	FBS	Clinical/chest radiographs/ sputum culture (LJ media)	DST (proportional methods)	New	21	6.66 (0.68–65.38)	ON
Fisher- Hoch SP et al. [50]	USA	S	1442 (401)	Not mentioned	Self-report	Sputum Culture	DST	All	21	2.14 (1.10–4.17)	Age, gender, alcohol and drug abuse, HIV infection, history of previous TB infection
Fisher- Hoch SP et al. [50] <sup>†</sup>	Mexico	S	1436 (287)	Not mentioned	Self-report	AFB/culture	DST	All	12	1.80 (1.13–2.87)	Age and gender
Gomez- Gomez A et al. [23]	Mexico	C	175 (56)	Record review	FBS/HbA1c	Culture/PCR for mycobacterium TB complex	DST (proportion method)	AII	T1 and T2	2.51 (1.11–5.67)	Age, sex, smoking history, chronic alcohol abuse, malnutrition, other illness conditions
Hafez S et al. [32]	Egypt	D	40 (12)	Face-to-face interview/ clinical/ laboratory	FBS	AFB/culture (U)	DST (proportional method)	All	T1 and T2	2.96 (0.73–11.93)	OZ
Hsu A et al. [48] <sup>*</sup>	Taiwan	S	139 (41)	Record review/ laboratory	FBS	Culture	DST (proportional method)	Previously treated	T1 and T2	1.52 (0.59–3.95)	Age, sex
Hsu A et al. [48]	Taiwan	S	869 (204)	Record review/ lab	FBS	Culture	DST (proportional method)	New	T1 and T2	0.95 (0.34–2.68)	Age, sex
Jitmuang A et al. [54]	Thailand	2	188 (31)	Record review	Not specified	Culture (LJ)	DST (proportional method)	AII	Type not mentioned	1.28 (0.54–3.02)	No
Magee MJ et al. [31] <sup>*</sup>	Peru	РС	823 (143)	Record review/ lab/interview	FBS/RBS/HbA1 c/medication with Insulin or Oral hypoglycemic agents	AFB	DST (Griess method)	New	T1 and T2	0.45 (0.26–0.78)	ON
Magee MJ et al. [53]	Georgia	Ы	263 (37)	Face-to-face interview/clinical	HbA1c¥	AFB/culture (LJ)/X-ray	DST (absolute concentration	New	Type not mentioned	2.27 (1.02–5.08)	Age, sex, HIV status and smoking

Magee MJ et al.     Feur     FC     848 (43)     Record review/ interview/Jab     Rest RBS/HbA1 by subino colona with by subino colo	IS/RBS/HbA1 AFB medication with sulin or Oral poglycemic agents b5 AFB/culture (LJ-media) cord review AFB/chest radiographs/ Culture b5 AFB/costitive culture b5 AFB/costitive chest X-ray/ clinical	method) DST (Griess method) DST (proportional method) DST (proportional method) DST (L1-media) NA NA DST (BACTEC, MGT)	Previously T1 and treated T2 New T2 Previously T2 treated T1 and All T1 and	T2 1.10 (0.56-2.19) 1.3 (0.6-2.8) 0.5 (0.2-1.1) T2 2.68 (1.05-6.86) (1.10-11.10) 3.50 3.50	No No Age and smoking
Magee M et al.     Peru     PC     848 (43)     Record review, lab.     Ensistion with insulin or Oral method)     DST (Griess method)     Pre       J11     Record review     FBS     AFB/culture     DST (groportional method)     Tet thod)       MIF et al. [49]     China     CS     192 (43)     Record review     FBS     AFB/culture     DST (groportional method)     Tet thod)       MIn J et al. [49]     China     CS     195 (55)     Record review     FBS     AFB/culture     DST (groportional method)     Tet thod)       Min J et al. [49]     Korea, Rep.     CC     195 (55)     Record review     FBS     AFB     DST (J-I-media)     AII       Min J et al. [52]     Korea, Rep.     CC     195 (55)     Record review     FBS     AFB     DST (J-I-media)     AII       Perez-Navaro LM     Mexico     CC     195 (55)     Record review     FBS     AFB     DST (J-I-media)     AII       Perez-Navaro LM     Mexico     CC     195 (55)     Record review     FBS     AFB     DST (J-I-media)     AII       Perez-	S/RBS/HbA1 AFB medication with medication with sulin or Oral poglycemic agents poglycemic agents AFB/culture (LJ-media) cord review AFB/culture (LJ-media) cord review AFB/culture S AFB AFB/culture S AFB AFB/contract AFB/contra	DST (Griess method) DST (proportional method) DST (proportional method) DST (LJ-media) NA NA DST (BACTEC, MGIT)	Previously T1 and treated T2 New T2 Previously T2 treated T1 and All T1 and	T2 1.10 (0.56-2.19) 1.3 (0.6-2.8) 0.5 (0.2-1.1) T2 2.68 (1.05-6.86) (1.05-6.86) (1.10-11.10) (1.10-11.10) 3.50	No No Age and smoking
Mi F et al. [49]*   China   CS   422 (144)   Record review   FBS   AFB/culture   DST (proportional)   Ne     Mi F et al. [49]   China   CS   199 (43)   Record review   FBS   AFB/culture   DST (proportional)   Pre     Min J et al. [52]   Korea, Rep.   CC   195 (55)   Record review   FBS   AFB/culture   DST (L-media)   All     Perez-Navarro LM   Mexico   CC   409 (146)   Record review   FBS   AFB   NA   All     Perez-Navarro LM   Mexico   CC   409 (146)   Record review   FBS   AFB   NA   All     Perez-Navarro LM   Mexico   PC   409 (146)   Record review   FBS   AFB   NA   All     Perez-Navarro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   NA   All     Perez-Navarro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   NA   All     Perez-Navarro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   NA   All<	S AFB/culture (LJ-media) (LJ-media) cord review AFB/culture (LJ-media) cond review AFB/culture Culture S AFB S AFB S AFB S AFB/positive chest X-ray/ clinical	DST (proportional method) DST (proportional method) DST (LJ-media) NA DST (BACTEC, MGIT)	New T2 Previously T2 treated All T1 and All T2	1.3 (0.6–2.8) 0.5 (0.2–1.1) (0.2–1.1) (0.2–1.1) (0.2–6.86) (1.05–6.86) (1.05–6.86) (1.10–11.10) 3.50	No No Age and smoking
Mir F et al. [49]   China   CS   199 (43)   Record review   FBS   AFB/culture   DST (proportional)   Prine     Min J et al. [52]   Korea, Rep.   CC   195 (55)   Record review   Record review   AFB/culture   DST (J-media)   All     Min J et al. [52]   Korea, Rep.   CC   195 (55)   Record review   Record review   AFB/culture   DST (J-media)   All     Prez-Navaro LM   Mexico   CC   409 (146)   Record review   FBS   AFB   NA   All     Prez-Navaro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   NA   All     Prez-Navaro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   NA   All     Prez-Navaro LM   Mexico   PC   507 (183)   Record review   FBS   AFB   MGIT)   All     Saktiawati AMI et   Indonesia   RC   356 (23)   Record review   FBS   AFB   MGIT)   All   All   MGIT)   All   All   All   All   All   All   All   All <td>S AFB/culture (LJ-media) cord review AFB/Chest radiographs/ Culture S AFB S AFB S AFB chest X-ray/ clinical</td> <td>DST (proportional method) DST (LJ-media) NA DST (BACTEC, MGIT)</td> <td>Previously T2 treated All T1 and All T2</td> <td>0.5 (0.2–1.1) (0.2–1.1) (1.05–6.86) (1.05–6.86) 3.50 (1.10–11.10) 3.50</td> <td>No Age and smoking Age</td>	S AFB/culture (LJ-media) cord review AFB/Chest radiographs/ Culture S AFB S AFB S AFB chest X-ray/ clinical	DST (proportional method) DST (LJ-media) NA DST (BACTEC, MGIT)	Previously T2 treated All T1 and All T2	0.5 (0.2–1.1) (0.2–1.1) (1.05–6.86) (1.05–6.86) 3.50 (1.10–11.10) 3.50	No Age and smoking Age
Min J et al. [52]   Korea, Rep.   CC   195 (55)   Record review   Record review   AFB/Chest   DST (J-media)   All     Perez-Navarro LM   Mexico   CC   409 (146)   Record review   EBS   AFB   NA   All     Perez-Navarro LM   Mexico   CC   409 (146)   Record review   EBS   AFB   NA   All     Perez-Navarro LM   Mexico   PC   507 (183)   Record review   EBS   AFB   NA   All     Saktiawati AMI et   Indonesia   RC   356 (23)   Record review   EBS   AFB/Ariay   All   All     al. [35]   Bangladesh   CC   1000 (83)   Face-to-face   Not specified   AFB/X-ray/   RIF/DST   All     Rifat M et al. [20]   Bangladesh   CC   1000 (83)   Face-to-face   Not specified   AFB/X-ray/   RIF/DST   All     Safita M et al. [20]   Bangladesh   CC   1000 (83)   Face-to-face   Not specified   AFB/X-ray/   RIF/DST   All     Safita M et al. [20]   Bangladesh   CC   1000 (83)   Face-to-face   Not sp	scord review AFB/Chest radiographs/ Culture S AFB AFB S AFB Culture Culture Culture AFB AFB AFB/positive Chest X-ray/ clinical	DST (LJ-media) NA DST (BACTEC, MGIT)	All T1 and All T2	T2 2.68 (1.05-6.86) 3.50 (1.10-11.10) 3.50	Age and smoking Age
Perez-Navarro LM     Mexico     CC     409 (146)     Record review     EBS     AFB     NA     All       et al. [19]     Perez-Navarro LM     Mexico     PC     507 (183)     Record review/     EBS     AFB     DST (BACTEC, Ne     Ne       Perez-Navarro LM     Mexico     PC     507 (183)     Record review/     EBS     AFB     MGIT)     Ne       Saktiawati AMI et al. [20]     Indonesia     RC     356 (23)     Record review     FBS     AFB/positive     GeneXpert (MTB/     All       al. [35]     Indonesia     RC     356 (23)     Record review     FBS     AFB/positive     GeneXpert (MTB/     All       al. [35]     Indonesia     RC     356 (23)     Record review     FBS     AFB/positive     GeneXpert (MTB/     All       al. [35]     Indonesia     RC     356 (23)     Record review     FBS     AFB/positive     GeneXpert (MTB/     All       al. [35]     Indonesia     RC     1000 (83)     Face-to-face     Not specified     ARMA-ray     MTB/RF/     MTB/RF/ <	S AFB S AFB/positive S AFB/positive chest X-ray/ clinical	NA DST (BACTEC, MGIT)	All T2	3.50 (1.10–11.10) 3.50	Age (
Perez-Navarro LM Mexico PC 507 (183) Record review/ self-report /lab EBS AFB DST (BACTEC, Ne Ne   Saktiawati AMI et Indonesia RC 356 (23) Record review FBS AFB/positive GeneXpert (MTB/ All   al. [35] Bangladesh C 356 (23) Record review FBS AFB/positive GeneXpert (MTB/ All   al. [35] Bangladesh C 1000 (83) Face-to-face Not specified AFB/X-ray/ clinical PCR (Xpert MTB/RF)/ record review All   Salindri A et al. Georgia PC 268 (36) Interview and record review Not specified AFB/X-ray PCR (Xpert MTB/RF)/ All   Salindri A et al. Georgia PC 268 (36) Interview/lab HbA1c Molecular DST (dasolute Ne   33] Fact At al. Georgia PC 268 (36) Interview/lab HbA1c Molecular DST (dasolute Ne	S AFB S AFB/positive chest X-ray/ clinical AFB/X-ray/	dst (Bactec, Mgit)		3.50	
Saktiawati AMI et Indonesia RC 356 (23) Record review FBS AFB/positive GeneXpert (MTB/ AII al. [35] al. [31] al. [31] al. [31] al. [35] al. [36] al. [36] al. [36] al. [37] al. [37] al. [38] al. [39] al. [39] al. [39] al. [39] al. [39] al. [30] al	S AFB/positive chest X-ray/ clinical AFR/V-ray		New T2	(1.60–7.10)	Age, sex, smoking, overcrowding
Rifat M et al. [20] Bangladesh CC 1000 (83) Face-to-face Not specified AFB/X-ray PCR (Xpert All   interview and interview and mTB/RIF/ mTB/RIF/ mTB/RIF/   record review record review culture/D5T/Line probe assay   Salindri A et al. Georgia PC 268 (36) Interview/lab HbA1c Molecular D5T (absolute Ne   [33] test/AFB/ method)	ot specified AFB/X-ray	GeneXpert (MTB/ RIF)/DST	All T2		
Salindri A et al. Georgia PC 268 (36) Interview/Iab HbA1c Molecular DST (absolute Ne diagnostic concentration test/AFB/ method)		PCR (Xpert MTB/RIF)/ culture/DST/Line probe assay	All T2	2.56 (1.51–4.34)	Age, education, occupation and smoking status
culture (LJ)/ clinical	DA1c Molecular diagnostic test/AFB/ culture (LJ)/ clinical	DST (absolute concentration method)	New Type n mentic	ot 2.51 ined (1.00–6.31)	Age, sex, education, income, smoking status, alcohol use, HIV status, kidney disease
Suarez-Garcia I et Spain CC 696 (41) Record review Not specified Culture DST (agar proportion All al. [21] method)	ot specified Culture	DST (agar proportion method)	All Type n mentic	ot 1.84 ined (0.53–6.33)	No
Tanrikulu A et al.     Turkey     CC     112 (9)     Record review     Not specified     Culture     DST (indirect     All       [28]     [28]     Proportion method)	ot specified Culture	DST (indirect proportion method)	All Type n mentic	ot 4.65 ined (1.01–21.51)	No
Zhang Q et al. China CS 2141 (203) Record review FBS AFB/sputum DST (proportional All [51] culture method)	S AFB/sputum	DST (proportional method)	All T1 and	T2 2.11 (1.42–3.11)	No

FMAC fine needle aspiration cytology, PCR polymerase chain reaction, AFB acid-fast bacilli, PC prospective cohort, RC retrospective cohort, CC case control, CS cross sectional, LI Lowenstein Jensen, HIV human immunodeficiency virus, TB-DM refers to the number of TB patients co-morbid with DM in the sample "These studies reported separate effect sizes for newly diagnosed and previously treated tuberculosis patients "This study reported separate effect sizes by country (USA and Mexico) "All effect sizes were presented as odd ratio

were cohort [14, 30, 31, 33–35, 53], and eight were cross-sectional [28, 48–51] by study design. No adjustment for confounding was done in 12 of the included studies, while the remaining studies adjusted for at least one confounding factor. The most commonly adjusted factors were age, sex, smoking status, and HIV status (Table 2). All the included studies passed quality assessment based on the Agency for Healthcare Research and Quality (AHRQ) [41] tool for cross-sectional studies and Newcastle-Ottawa Scale (NOS) [40] criteria for case-control and cohort studies (Additional file 2: Table S2).

## Associations between DM and MDR-TB

We explored the influence of each individual study on the overall meta-analysis summary estimate. Table 3 shows the influence of omitting a single study on the overall summary estimate. Accordingly, we identified that omitting Magee MJ et al. [31] resulted in a large improvement on the overall summary estimate compared to the combined summary estimate obtained by omitting any one single study included in this meta-analysis (Table 3). This study was then excluded from the rest of the analysis resulting in 24 observational studies considered to conduct this meta-analysis.

Based on the meta-analysis of 24 observational studies, the overall pooled effect estimate was 1.97 (OR = 1.97, 95% CI 1.58 to 2.45,  $I^2 = 38.2\%$ , *P* value for heterogeneity = 0.031). This result indicated a 97% increased risk of MDR-TB among TB-DM co-morbid patients. The significant and positively increased risk of MDR-TB among TB-DM co-morbid patients remained in the same direction in a subgroup analysis by study characteristics. When seen by level of adjustment to a confounding factor, studies which adjusted for at least one confounding factor had more than a two-fold increased risk of MDR-TB (OR = 2.43, 95% CI 1.90 to 3.12) (Fig. 2). In a subgroup analysis of studies which adjusted for a minimum of one confounding factor, a strong positive association was observed by study design (cross-sectional, 4 studies, OR = 1.72, 95% CI 1.23 to 2.41; case-control, 5 studies, OR = 2.89, 95% CI 2.02 to 4.12; cohort, 4 studies, OR = 3.36, 95% CI 1.82 to 6.20) and method of TB diagnosis (culture confirmed, 8 studies, OR = 1.97, 95% CI 1.51 to 2.57; Sputum smear test, 5 studies, OR = 3.73 95% CI 2.33 to 5.97) (Table 4).

To evaluate consistency of the evidence over the years along with its sufficiency, we ran a cumulative meta-analysis which calculates effect estimates as newer studies are added. Accordingly, between 2001 and 2014, there was a positive association between DM and MDR-TB but with a swinging effect size. However, with the addition of three case-control and four cohort studies between 2014 and 2018, a strong evidence that

Table 3 Single study influence analysis on the overall meta-
analysis summary estimate of the association between diabete
mellitus and multi-drug-resistant tuberculosis

Study omitted	Effect size,	95% CI	
	OR	LCL	UCL
Gomez-Gomez A et al. [23]	1.81	1.37	2.38
Magee MJ et al. [53]	1.82	1.38	2.40
Fisher-Hoch SP et al. <sup>†</sup> [50]	1.84	1.38	2.45
Hsu A et al.* [48]	1.85	1.40	2.44
Saktiawati AMI et al. [35]	1.75	1.36	2.26
Rifat M et al. [20]	1.80	1.36	2.38
Bashar M et al. [18]	1.76	1.35	2.29
Perez-Navarro LM et al. [30]	1.77	1.35	2.33
Fisher-Hoch SP et al. [50]	1.82	1.38	2.41
Perez-Navarro LM et al. [19]	1.80	1.37	2.35
Salindri AD et al. [33]	1.81	1.38	2.38
Min J et al. [52]	1.81	1.37	2.38
Hsu A et al. [48]	1.88	1.43	2.46
Chang JT [14]	1.83	1.38	2.42
Zhang Q et al. [51]	1.82	1.37	2.43
Hafez S et al. [32]	1.81	1.38	2.38
Magee MJ et al.* [31]	1.88	1.43	2.48
Jitmuang A et al. [54]	1.86	1.41	2.46
Mi F et al.* [49]	1.93	1.49	2.49
Suarez-Garcia I et al. [21]	1.83	1.40	2.41
Magee MJ et al. [31]	1.97	1.58	2.45
Baghaei P et al. [25]	1.88	1.44	2.46
Tanrikulu A et al. [28]	1.79	1.37	2.35
Mi F et al. [49]	1.87	1.41	2.46
Carreira S et al. [34]	1.84	1.40	2.41
Combined	1.83	1.40	2.39

UCL upper confidence limit, LCL lower confidence limit, Cl confidence interval, OR odds ratio

\*These studies reported separate effect sizes for newly diagnosed and previously treated tuberculosis patients

This study reported separate effect sizes by country (USA and Mexico)

TB patients co-morbid with DM had an increased risk of developing MDR-TB has sustained (Fig. 3).

There was no significant publication bias found either by the Egger's regression asymmetry test or by a funnel plot (Fig. 4a). The contour-enhanced funnel plot examination (Fig. 4b) confirms this, which distinguished between publication bias and other causes. It showed that small studies were found not only in the areas of statistical significance but also in areas of non-statistical significance.

## Discussion

In this study, we pooled estimates on the association between DM and MDR-TB from 24 observational studies and identified a significant and positive association

AuthorYr	Design	Population		OR (95% CI)
Adjusted at least for one factor				
Gomez-Gomez A et al.(2015)	Case control	175		2.51 (1.11, 5.67)
Magee MJ et al.(2015)	Prospective cohort	263	•	2.27 (1.02, 5.08)
Fisher-Hoch SP et al.(2008)	Cross-sectional	1436		1.80 (1.13, 2.87)
Hsu A et al.(2012)	Cross-sectional	139		1.52 (0.59, 3.95)
Saktiawati AMI et.al(2018)	Retrospective cohort	356		• 17.90 (3.30, 96.80)
Rifat M et al.(2014)	Case control	1000		2.56 (1.51, 4.34)
Bashar M et al.(2001)	Case control	155	• • • • • • • • • • • • • • • • • • •	5.30 (1.90, 14.70)
Perez-Navarro LM et al.(2017)	Prospective cohort	507	<b>·</b>	3.50 (1.60, 7.10)
Fisher-Hoch SP et al.(2008)	Cross-sectional	1442		2.14 (1.10, 4.17)
Perez-Navarro LM et al.(2015)	Case control	409		3.50 (1.10, 11.10)
Salindri AD et al.(2016)	Prospective cohort	268	•	2.51 (1.00, 6.31)
Min J et al.(2005)	Case control	195	•	2.68 (1.05, 6.86)
Hsu A et al.(2012)	Cross-sectional	869	-	0.95 (0.34, 2.68)
Subtotal (I-squared = 19.1%, p = 0.250)			$\diamond$	2.43 (1.90, 3.12)
No adjustment				
Chang JT(2011)	Prospective cohort	192	•	1.95 (1.00, 3.81)
Zhang Q et al.(2009)	Cross-sectional	2141		2.11 (1.42, 3.11)
Hafez S et al.(2013)	Case control	40	•	2.96 (0.73, 11.93)
Magee MJ et al.(2013)	Prospective cohort	848		1.10 (0.56, 2.19)
Jitmuang A et al.(2015)	Case control	188	<u> </u>	1.28 (0.54, 3.02)
Mi F et al.(2014)	Cross-sectional	199	┝╼╋╴╎	0.50 (0.20, 1.10)
Suarez-Garcia I et al.(2009)	Case control	696	• •	1.84 (0.53, 6.33)
Baghaei P et al.(2009)	Case control	282	•	0.68 (0.19, 2.36)
Tanrikulu A et al.(2008)	Cross-sectional	112		4.65 (1.01, 21.51)
Mi F et al.(2014)	Cross-sectional	422	<b>—</b>	1.30 (0.60, 2.80)
Carreira S et al.(2012)	Retrospective cohort	246		1.49 (0.24, 9.06)
Subtotal (I-squared = 35.4%, p = 0.115)			$\diamond$	1.45 (1.05, 2.02)
Overall (I-squared = 38.2%, p = 0.031)			\$	1.97 (1.58, 2.45)
NOTE: Weights are from random effects ana	lysis			
		ا د	1 1	-
		Reducerisk of MDR-TB	I Increase risk of M	DR-TB
		Odd	ls Ratio	

with a low between-study heterogeneity. In a pooled estimate of studies which adjusted for at least one confounding factor, stronger association was observed by study design, type of TB, method of TB diagnosis, type of DM, method of DM diagnosis, and country income level.

The pooled effect estimate, in subgroup analyses of 13 studies which adjusted for at least one confounding factor, on the association between DM and MDR-TB (OR = 2.43, 95% CI 1.90 to 3.12,  $I^2 = 19.1\%$ , P value = 0.250) was consistent in the direction of association, but stronger than what was reported in a previous systematic review and meta-analysis (OR = 1.71, 95% CI 1.32 to 2.22,  $I^2 = 46.8\%$ , P value = 0.020) [29]. This difference could be due to the fact that Liu et al. [29] mixed three studies which reported unadjusted OR [21, 49, 54], while we included three new cohort studies that reported adjusted OR [30, 33, 35]. Compared to a similar prior meta-analysis [29], we included 10 additional studies [14, 25, 28, 30-35, 51] and supplemented this with strong analytical rigor. Thus, our meta-analysis fortifies the evidence base for the association between DM and MDR-TB. Our results are also in agreement with a review that showed TB and DM co-morbidity were more likely to be evident among people with anti-TB drug resistance [55], signifying a clear association between the two diseases. In another review, Baker et al. [8] did not find an increased risk of MDR-TB among people with DM. However, it is worth mentioning that this study aimed to review literature on the impact of DM on TB outcomes, and this particular sub-analysis included only four studies. Although not having MDR-TB as an outcome, another review noted significant associations in the pooled risk of association

Study characteristics	Studies that ad	justed for at least one c	ovariate		Studies that dic	I not adjustment for co	variates	
	No. of studies	Pooled OR (95% CI)	l <sup>2</sup> (%)	P value <sup>¥</sup>	No. of studies	Pooled OR (95% CI)	l <sup>2</sup> (%)	P value <sup>¥</sup>
Study design								
Case control	5	2.89 (2.02, 4.12)	0.0	0.773	4	1.39 (0.79, 2.45)	0.0	0.459
Cohort	4	3.36 (1.82, 6.20)	40.2	0.171	3	1.47 (0.93, 2.34)	0.0	0.501
Cross sectional	4	1.72 (1.23, 2,41)	19.1	0.250	4	1.45 (0.68, 3.09)	73.2	0.011
Type of TB								
Both or not defined	7	2.72 (1.88, 3.94)	38.4	0.136	7	1.90 (1.40, 2.60)	0.0	0.492
New	5	2.36 (1.59, 3.51)	3.8	0.385	2	1.64 (0.99, 2.71)	0.0	0.436
Previously treated	1 <sup>+</sup>				2	0.77 (0.36, 1.67)	35.4	0.115
Diagnosis of TB								
Culture confirmed	8	1.97 (1.51, 2.57)	0.0	0.831	8	1.60 (1.08, 2.38)	43.6	0.088
Sputum smear test only	5	3.73 (2.33, 5.97)	28.0	0.235	3	1.03 (0.58, 182)	0.0	0.736
Type of DM								
Type 2	6	2.67 (1.82, 3.93)	40.7	0.134	3	1.12 (0.52, 2.42)	67.3	0.047
Type 1 and type 2	5	2.22 (1.31, 3.76)	35.8	0.183	3	1.77 (1.08, 2.89)	35.5	0.212
Not defined	2	2.37 (1.29, 4.34)	0.0	0.872	5	1.46 (0.84, 2.51)	0.0	0.428
Diagnosis of DM								
HbA1c or FBS	3	2.42 (1.49, 3.93)	0.0	0.981	1 <sup>+</sup>			
Only FBS	6	2.68 (1.45, 4.96)	53.5	0.056	6	1.50 (0.93, 2.42)	51.6	0.066
Self-report/unspecified	4	2.33 (1.65, 3.29)	21.5	0.282	4	1.49 (0.77, 2.89)	21.9	0.279
Country income level*								
Lower middle income	4	2.95 (1.69, 5.16)	40.6	0.168	1 <sup>+</sup>			
Upper middle income	4	2.32 (1.65, 3.28)	0.0	0.418	7	1.27 (0.80, 2.01)	56.9	0.031
High income	5	2.13 (1.29, 3.52)	34.6	0.191	3	1.88 (1.07, 3.29)	0.0	0.963
Overall	13	2.43 (1.90, 3.12)	19.1	0.250	11	1.45 (1.05, 2.02)	35.4	0.115

**Table 4** Subgroup analyses of 24 observational studies on the association between diabetes mellitus and multi-drug-resistant-tuberculosis among tuberculosis patients co-morbid with diabetes mellitus

OR odds ratio,  $l^2$  the variation in estimate attributable to heterogeneity, ¥ P value for heterogeneity, Cl confidence interval, TB tuberculosis, DM diabetes mellitus, FBS fasting blood sugar, HbA1c glycosylated hemoglobin

<sup>†</sup>Effect estimate not pooled due limited number of available studies

\*Income level based on World Bank Classification (https://datahelpdesk.worldbank.org/knowledgebase/articles/906519-world-bank-country-and-lending-groups)

between DM and active TB without regarding the type of study design implemented [56].

Signifying a stability of association, we found a significant and positive association between DM and MDR-TB in subgroup analyses of studies which adjusted for at least one confounding factor. The association persisted irrespective of the country where the primary studies were conducted which covered diverse population groups, how DM or TB diagnosis was made, and study design of the primary studies. However, pooled estimate of a weaker strength was found in studies which did not adjust for at least one confounding factor (OR = 1.45, 95% CI 1.05 to 2.02). We included studies which reported unadjusted effect estimate to reduce inflation of the pooled effect estimate from the adjusted only analysis, increase the number of available studies for analysis, and widen the representativeness of our findings [57]. Although unadjusted findings lack adjustment by statistical methods, they attempted to adjust at the design level albeit still suffering from an unobserved confounding effect. Therefore, we suggest our findings should be interpreted with caution.

In subgroup analyses of studies which adjusted for at least one confounding factor on the association between DM and MDR-TB, there was an increased risk in lower middle-income countries (OR = 2.95, 95% CI 1.69 to 5.16) compared to that found in upper middle-income countries (OR = 2.32, 95% CI 1.65 to 3.28) and high-income countries (OR = 2.13, 95% CI 1.29 to 3.52). This might be due to more studies in high-income- or upper middle-income country settings diagnosed DM by blood test (FBS or HbA1c) or confirmed TB diagnosis based on culture which may have reduced bias due to misclassification [19, 23, 30, 48, 50, 52]. Consistent with our

AuthorYr	Design		OR (95% CI)
Bashar M et al.(2001)	Case control		5.30 (1.91, 14.74)
Min J et al.(2005)	Case control	│   — ◆ —	- 3.66 (1.83, 7.31)
Fisher-Hoch SP et al.(2008)	Cross-sectional	│ <u>→</u>	2.79 (1.70, 4.60)
Fisher-Hoch SP et al.(2008)	Cross-sectional		2.31 (1.56, 3.41)
Tanrikulu A et al.(2008)	Cross-sectional	│	2.38 (1.65, 3.44)
Suarez-Garcia I et al.(2009)	Case control		2.26 (1.65, 3.10)
Zhang Q et al.(2009)	Cross-sectional		2.20 (1.72, 2.82)
Baghaei P et al.(2009)	Case control	│	2.13 (1.62, 2.80)
Chang JT(2011)	Prospective cohort	<b>→</b>	2.09 (1.66, 2.63)
Hsu A et al.(2012)	Cross-sectional	│ <b>→</b>	2.02 (1.57, 2.58)
Carreira S et al.(2012)	Retrospective cohort	│	2.01 (1.60, 2.52)
Hsu A et al.(2012)	Cross-sectional		1.98 (1.60, 2.45)
Magee MJ et al.(2013)	Prospective cohort	│	1.87 (1.49, 2.33)
Hafez S et al.(2013)	Case control		1.89 (1.53, 2.34)
Rifat M et al.(2014)	Case control		1.96 (1.61, 2.39)
Mi F et al.(2014)	Cross-sectional	│	1.92 (1.58, 2.32)
Mi F et al.(2014)	Cross-sectional		1.76 (1.38, 2.24)
Perez-Navarro LM et al.(2015)	Case control	│	1.80 (1.42, 2.29)
Jitmuang A et al.(2015)	Case control		1.77 (1.41, 2.23)
Gomez-Gomez A et al.(2015)	Case control		1.80 (1.45, 2.25)
Magee MJ et al.(2015)	Prospective cohort		1.83 (1.48, 2.25)
Salindri AD et al.(2016)	Prospective cohort	│	1.85 (1.51, 2.26)
Perez-Navarro LM et al.(2017)	Prospective cohort		1.91 (1.56, 2.33)
Saktiawati AMI et.al(2018)	Retrospective cohort		1.97 (1.58, 2.45)
	.1	1 I	1 I 10 15
	Reducerisk of ME	DR-TB Increase risk of MD	R-TB
		Odds Ratio	

finding that there was an increased risk of MDR-TB among TB-DM co-morbid patients in lower middle-income countries, a meta-analysis by Al-Rifai et al. [56] reported that there was a higher risk of TB-DM association in lower middle-income countries compared to high-and upper middle-income countries.

A strong association was found in subgroup analyses by TB type in studies which reported adjusted estimate, where DM increased the risk of MDR-TB among new TB patients (OR = 2.36, 95% CI 1.59 to 3.51) and in those with undefined TB type (OR = 2.72, 95% CI 1.88 to 3.94). However, due to inadequate number of studies which adjusted for at least one confounding factor and included only previously treated TB patients, we did not present a pooled estimate on the risk of MDR-TB among previously treated TB-DM co-morbid patients. Therefore, we are not able to confirm whether DM is indeed not a risk factor for MDR-TB among previously treated TB patients. The meta-analysis by Liu et al. [29] also reported similar findings on the risk of any MDR-TB (OR = 1.94, 95% CI 1.42 to 2.65) and primary MDR-TB (OR = 1.69, 95% CI 1.09 to 2.62) among TB patients co-morbid with DM. Due to the addition of three cohort studies [30, 33, 35], which were not included in the Liu et al.'s meta-analysis [29], we reported a stronger pooled effect estimate for the increase in the risk of MDR-TB among new- and any-TB patients co-morbid with DM. Similarly, irrespective of the study design used in the primary studies which adjusted for at least one confounding factor, we consistently found a significant and positive association between DM and MDR-TB in cross-sectional (OR = 1.72 95% CI 1.23 to 2.41), case-control (OR = 2.89, 95% 2.02 to 4.12), and cohort studies (OR = 3.36, 95% CI 1.82, 6.20). Similar findings were reported by Liu et al. for case-control studies but not for cross-sectional studies [29]. Because only one



cohort study was included [53] by Liu et al. [29], we cannot compare the pooled estimate from cohort studies in our meta-analysis to theirs. Though it is difficult to establish causality based on evidence from observational studies, we believe that these pooled estimates from studies which controlled for potential confounding factors further solidify the existence of a strong association between DM and MDR-TB among patients co-morbid with TB.

In this study, we have shown the most substantial evidence to date on the association between DM and MDR-TB by including a comprehensive list of studies carried out in various settings around the world. Methodologically, we conducted a cumulative meta-analysis to see the trend of evidence and performed a single study influence analysis and subgroup analysis. However, we would also like to acknowledge potential limitations of this study. Firstly, we have not included studies published in non-English languages. Given majority of MDR-TB cases and co-morbidities are from Russia, China, and India [1], exclusion of studies from these countries might bias our finding. Secondly, the inclusion of studies which did not consistently define the type of TB and did not control for potential confounding factors increased the observed heterogeneity among studies. However, in an attempt to deal with this, we have performed and reported sensitivity analysis. Finally, potential misclassification regarding a non-uniform diagnosis of DM and TB among participants included in the different studies might have contributed to varying risk estimates. This is particularly observed in the stratified analysis based on diagnosis of DM and TB, where higher estimates were observed for those diagnosed based on information obtained from medical records and selfreports, (diagnosis of DM: FBS from medical record, OR = 2.68, 95% CI 1.45 to 4.96; HbA1c, OR = 2.42, 95% CI 1.49 to 3.93; self-report, OR = 2.33, 95% CI 1.65 to 3.29) and (diagnosis of TB: culture confirmed, OR = 1.97, 95% CI 1.51 o 2.57; AFB only from medical record, OR = 3.73, 95% CI 2.33 to 5.97).

In conclusion, our results demonstrate that a more robust TB treatment and follow-up might be necessary for patients with DM. In light of the global DM epidemic [5], this study emphasizes the message that there is a strong need for a bi-directional screening and co-management approach in the attempt to halt the TB-DM co-morbidity [58]. Even though countries representing diverse income levels were not included in our study, we stress the need to maximize efforts to prevent DM and TB co-morbidity and reduce the burden of MDR-TB in countries with varying economical standings. The integrated and collaborative effort between TB and DM control programs will ultimately count on achieving the global "End TB Strategy" [2]. Efforts to control DM can have a substantial beneficial effect on TB outcomes, particularly in the case of MDR-TB. Policymakers can focus on new targets pertaining to an enhanced care plan for DM patients with TB, particularly among the slightest evidence of problems with adherence or prolonged and complicated infections. Furthermore, focus on the early identification and treatment of individuals with the co-morbidity can result in an enhanced treatment outcome. We recommend future prospective cohort studies to focus on bacteriologically confirmed TB cases that objectively diagnose DM, with clearly defined types of both TB and DM coupled with robust controls for potential confounding.

## **Additional files**

Additional file 1: Table S1. Preferred Reporting Items for Systematic Reviews and Meta-analysis (PRISMA) 2009 checklist (DOCX 67 kb)

Additional file 2: Table S2. Study quality assessment results for case-control, cohort, and cross-sectional studies (DOCX 26 kb)

Additional file 3: List of studies excluded with reasons after full-text review (DOCX 20 kb)

#### Abbreviations

AFB: Acid-fast bacilli; CI: Confidence interval; DM: Diabetes mellitus; DST: Drug susceptibility test; FBS: Fasting blood sugar; FNAC: Fine needle aspiration; HIV: Human immunodeficiency virus; LJ: Lowenstein Jensen; MDR-TB: Multi-drug-resistant tuberculosis; OR: Odds ratio; PCR: Polymerase chain reaction; TB: Tuberculosis; WHO: World Health Organization

#### Acknowledgements

We gratefully acknowledge Sjoukje van der Werf (Medical Information Specialist, University Medical Center Groningen, Groningen, The Netherlands) for her invaluable support in the development of search strings.

## Availability of data and materials

All data pertaining to the findings are presented in this paper. However, the data can be obtained from the corresponding author at any time on request.

## Authors' contributions

BS and TD conceived and designed the study. BS and TD developed the search strings. BS, MM, and TD involved in screening, data extraction, and quality assessment of studies. BS, MM, and MA analyzed and interpreted findings. BS, MM, TD, AT, and MA wrote the manuscript. All of the authors provided critical comments for revision and approved the final version of the manuscript.

## Ethics approval and consent to participate

Not applicable

#### Consent for publication

Not applicable

#### **Competing interests**

The authors declare that they have no competing interests.

## **Publisher's Note**

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

#### Author details

<sup>1</sup>Department of Epidemiology and Biostatistics, College of Health and Medical Sciences, Haramaya University, Harar, Ethiopia. <sup>2</sup>Department of Epidemiology, University of Groningen, University Medical Center Groningen, Groningen, The Netherlands. <sup>3</sup>Division of Epidemiology, College of Public Health, The Ohio State University, Columbus, OH, USA. <sup>4</sup>Amref Health Africa in Ethiopia, Monitoring, Evaluation and Research Unit, Addis Ababa, Ethiopia.

## Received: 21 June 2018 Accepted: 1 October 2018 Published online: 15 October 2018

#### References

- World Health Organization. Global tuberculosis report 2017. Geneva: World Health Organization; 2017. [cited 2018 Apr 20]. Available from: http://www. who.int/tb/publications/global\_report/en/
- World Health Organization. WHO end TB strategy: global strategy and targets for tuberculosis prevention, care and control after 2015. Geneva: WHO; 2015. [cited 2018 April 30]. Available from: http://www.who.int/tb/ post2015\_strategy/en/
- Lönnroth K, Roglic G, Harries AD. Improving tuberculosis prevention and care through addressing the global diabetes epidemic: from evidence to policy and practice. Lancet Diabetes Endocrinol. 2014;2:10.

- Marais BJ, Lönnroth K, Lawn SD, Migliori GB, Mwaba P, Glaziou P, et al. Tuberculosis comorbidity with communicable and non-communicable diseases: integrating health services and control eff orts. Lancet Infect Dis 2013, 2013;13:13.
- IDF. IDF Diabetes ATLAS 2013 [cited 2018 Apr 27]. 6th:[Available from: www. idf.org/diabetesatlas.
- Alkabab YM, Al-Abdely HM, Heysell SK. Diabetes-related tuberculosis in the Middle East: an urgent need for regional research. Int J Infect Dis. 2015;40:7.
- Jeon CY, Murray MB. Diabetes mellitus increases the risk of active tuberculosis: a systematic review of 13 observational studies. PLoS Med. 2008;5:e152.
- Baker MA, Harries AD, Jeon CY, Hart JE, Kapur A, Lönnroth K, et al. The impact of diabetes on tuberculosis treatment outcomes: a systematic review. BMC Med. 2011;9:81.
- World Health Organization. World Health Organization multidrug and extensively drug-resistant TB (M/XDR-TB): 2010 global report on surveillance and response. Geneva: WorldHealthOrganization; 2010.
- 10. Nations JA, Lazarus AA, Walsh TE. Drug-resistant tuberculosis. Dis Mon. 2006;52:435–40.
- 11. Jain A, Dixit P. Multidrug resistant to extensively drug resistant tuberculosis: what is next? J Biosci. 2008;33:17.
- 12. Ormerod LP. Multidrug-resistant tuberculosis (MDR-TB): epidemiology, prevention and treatment. Br Med Bull. 2005;73-74:8.
- Kang YA, Kim SY, Jo KW, Kim HJ, Park SK, Kim TH, et al. Impact of diabetes on treatment outcomes and long-term survival in multidrug-resistant tuberculosis. Respiration. 2013;86:7.
- Chang J, Dou H, Yen C, Wu Y, Huang R, Lin H, et al. Effect of type 2 diabetes mellitus on the clinical severity and treatment outcome in patients with pulmonary tuberculosis: a potential role in the emergence of multidrug-resistance. J Formosan Med Assoc. 2011;110:10.
- Kameda K, Kawabata S, Masuda N. Follow-up study of short course chemotherapy of pulmonary tuberculosis complicated with diabetes mellitus. Kekkaku. 1990;65:13.
- Park S, Shin J, Kim J, Park I, Choi B, Choi J, et al. The effect of diabetic control status on the clinical features of pulmonary tuberculosis. Eur J Clin Microbiol Infect Dis. 2012;31:6.
- Baghaei P, Marjani M, Javanmard P, Tabarsi P, Masjedi MR. Diabetes mellitus and tuberculosis facts and controversies. J Diabetes Metab Disord. 2013;12(1):58.
- Bashar M, Alcabes P, Rom WN, Condos R. Increased incidence of multidrugresistant tuberculosis in diabetic patients on the Bellevue Chest Service, 1987 to 1997. CHEST Journal. 2001;120:6.
- Pérez-Navarro LM, Fuentes-Domínguez FJ, Zenteno-Cuevas R. Type 2 diabetes mellitus and its influence in the development of multidrug resistance tuberculosis in patients from southeastern Mexico. J Diabetes Complicat. 2015;29:6.
- Rifat M, Milton AH, Hall J, Oldmeadow C, Islam MA, Husain A, et al. Development of multidrug resistant tuberculosis in Bangladesh: a casecontrol study on risk factors. PLoS One. 2014;9:e105214.
- Suarez-Garcia I, Rodriguez-Blanco A, Vidal-Perez J, Garcia-Viejo M, Jaras-Hernandez M, Lopez O, et al. Risk factors for multidrug-resistant tuberculosis in a tuberculosis unit in Madrid, Spain. Eur J Clin Microbiol Infect Dis. 2009;28:6.
- 22. Bendayan D, Hendler A, Polansky V, Weinberger M. Outcome of hospitalized MDR-TB patients: Israel 2000–2005. Eur J Clin Microbiol Infect Dis. 2011;30:5.
- 23. Gómez-Gómez A, Magaña-Aquino M, López-Meza S, Aranda-Álvarez M, Díaz-Ornelas DE, Hernández-Segura MG, et al. Diabetes and other risk factors for multi-drug resistant tuberculosis in a Mexican population with pulmonary tuberculosis: case control study. Arch Med Res. 2015;46:7.
- Kikvidze M, Mikiashvili L. Impact of diabetes mellitus on drug-resistant tuberculosis treatment outcomes in Georgia-cohort study. Eur Respir J. 2013;42.
- Baghaei P, Tabarsi P, Chitsaz E, Novin A, Alipanah N, Kazempour M, et al. Risk factors associated with multidrug-resistant tuberculosis. TANAFFOS-Journal of Respiratory Disease, Thoracic Surgery Intensive Care and Tuberculosis. 2009;8:5.
- Baghaei P, Tabarsi P, Abrishami Z, Mirsaeidi M, Faghani YA, Mansouri SD, et al. Comparison of pulmonary TB patients with and without diabetes mellitus type II. Tanaffos. 2010;9:8.
- 27. Singla R, Khan N, Al-Sharif N, Al-Sayegh M, Shaikh M, Osman M. Influence of diabetes on manifestations and treatment outcome of pulmonary TB

patients. The International Journal of Tuberculosis and Lung Disease. 2006;10:6.

- 28. Tanrikulu AC, Hosoglu S, Ozekinci T, Abakay A, Gurkan F. Risk factors for drug resistant tuberculosis in southeast Turkey. Trop Dr. 2008;38:4.
- Liu Q, Li W, Xue M, Chen Y, Du X, Wang C, et al. Diabetes mellitus and the risk of multidrug resistant tuberculosis: a meta-analysis. Sci Rep. 2017;7(1):1090.
- Perez-Navarro LM, Restrepo BJ, Fuentes-Dominguez FJ, Duggirala R, Morales-Romero J, López-Alvarenga JC, et al. The effect size of type 2 diabetes mellitus on tuberculosis drug resistance and adverse treatment outcomes. Tuberculosis. 2017;103:9.
- Magee M, Bloss E, Shin S, Contreras C, Huaman HA, Ticona JC, et al. Clinical characteristics, drug resistance, and treatment outcomes among tuberculosis patients with diabetes in Peru. Int J Infect Dis. 2013;17:e404–e12.
- Hafez S, Elhefnawy A, Hatata E, El Ganady A, Ibrahiem M. Detection of extensively drug resistant pulmonary tuberculosis. Egyptian Journal of Chest Diseases and Tuberculosis. 2013;62:12.
- Salindri AD, Kipiani M, Kempker RR, Gandhi NR, Darchia L, Tukvadze N, et al. Diabetes reduces the rate of sputum culture conversion in patients with newly diagnosed multidrug-resistant tuberculosis. Open Forum Infect Dis. 2016;3(3):ofw126.
- Carreira S, Costeira J, Gomes C, André J, Diogo N. Impact of diabetes on the presenting features of tuberculosis in hospitalized patients. Rev Port Pneumol (English Edition). 2012;18((5)):239–43.
- Saktiawati AMI, Subronto YW. Influence of diabetes mellitus on the development of multi-drug resistant-tuberculosis in Yogyakarta. Indones J Intern Med. 2018;50(1):7.
- Tegegne BS, Habtewold TD, Mengesha MM, Burgerhof JG. Association between diabetes mellitus and multi-drug-resistant tuberculosis: a protocol for a systematic review and meta-analysis. Syst Rev. 2017;6:6.
- Beller EM, Glasziou PP, Altman DG, Hopewell S, Bastian H, Chalmers I, et al. PRISMA for abstracts: reporting systematic reviews in journal and conference abstracts. PLoS Med. 2013;10:e1001419.
- Moher D, Schulz KF, Simera I, Altman DG. Guidance for developers of health research reporting guidelines. PLoS Med. 2010;7:e1000217.
- Stroup DF, Berlin JA, Morton SC, Olkin I, Williamson GD, Rennie D. Metaanalysis of observational studies in epidemiology. A Proposal for Reporting JAMA. 2000;283(15):5.
- Wells G, Shea B, O'connell D, Peterson J, Welch V, Losos M, et al. The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses 2000 [cited 2016 August 20]. Available from: http:// www.ohri.ca/programs/clinical\_epidemiology/oxford.htm.
- Zeng X, Zhang Y, Kwong JS, Zhang C, Li S, Sun F, et al. The methodological quality assessment tools for preclinical and clinical studies, systematic review and meta-analysis, and clinical practice guideline: a systematic review. J Evid Based Med. 2015;8:9.
- 42. Cochran WG. The combination of estimates from different experiments. Biometrics. 1954;10:30.
- Higgins J, Thompson SG. Quantifying heterogeneity in a meta-analysis. Stat Med. 2002;21:30.
- 44. DerSimonian R, Laird N. Meta-analysis in clinical trials. Control Clin Trials. 1986;7:11.
- Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. BMJ. 1997;315:5.
- Tob'ıas A. sbe26. Assessing the influence of a single study in the metaanalysis estimate. Stata Tech Bull. 1999;47:3.
- StataCorp. Stata Statistical Software: Release 14. College Station: StataCorp LP; 2015.
- Hsu A, Lee J, Chiang C, Li Y, Chen L, Lin C. Diabetes is associated with drugresistant tuberculosis in Eastern Taiwan. Int J Tuberc Lung Dis. 2013;17:13.
- Mi F, Jiang G, Du J, Li L, Yue W, Harries AD, et al. Is resistance to antituberculosis drugs associated with type 2 diabetes mellitus? A register review in Beijing, China. Glob Health Action. 2014;7:24022.
- Fisher-Hoch SP, Whitney E, McCormick JB, Crespo G, Smith B, Rahbar MH, et al. Type 2 diabetes and multidrug-resistant tuberculosis. Scand J Infect Dis. 2008;40:6.
- 51. Zhang Q, Xiao H, Sugawara I. Tuberculosis complicated by diabetes mellitus at Shanghai Pulmonary Hospital. China Jpn J Infect Dis. 2009;62:12.
- 52. Min J, Park K, Whang S, Kim J. Risk factors for primary multidrug resistant tuberculosis. Tuber Res Dis. 2005;59(6).

- Magee MJ, Kempker RR, Kipiani M, Gandhi NR, Darchia L, ea TN. Diabetes mellitus is associated with cavities, smear grade, and multidrug-resistant tuberculosis in Georgia. Int J Tuberc Lung Dis. 2015;19:8.
- Jitmuang A, Munjit P, Foonglada S. Prevalence and factors associated with multidrug-resistant tuberculosis at Siriraj Hospital, Bangkok, Thailand. Southeast Asian J Trop Med Public Health. 2015;46:10.
- Workneh MH, Bjune GA, Yimer SA. Prevalence and associated factors of tuberculosis and diabetes mellitus comorbidity: a systematic review. PLoS One. 2017;12:e0175925.
- Al-Rifai RH, Pearson F, Critchley JA, Abu-Raddad LJ. Association between diabetes mellitus and active tuberculosis: a systematic review and metaanalysis. PLoS One. 2017;12:e0187967.
- Voils CI, Crandell JL, Chang Y, Leeman J, Sandelowski M. Combining adjusted and unadjusted findings in mixed research synthesis. J Eval Clin Pract. 2011;17(3):10.
- World Health Organization, and International Union against Tuberculosis and Lung Disease. Collaborative framework for care and control of tuberculosis and diabetes. Geneva: World Health Organization; 2011. [cited 2018 May 12]. Available from: http://www.who.int/iris/handle/10665/44698

#### Ready to submit your research? Choose BMC and benefit from:

- fast, convenient online submission
- thorough peer review by experienced researchers in your field
- rapid publication on acceptance
- support for research data, including large and complex data types
- gold Open Access which fosters wider collaboration and increased citations
- maximum visibility for your research: over 100M website views per year

#### At BMC, research is always in progress.

Learn more biomedcentral.com/submissions

